The Burden of the Nondiversifiable Risk of Entrepreneurship

BY ROBERT E. HALL AND SUSAN E. WOODWARD* 

Entrepreneurship is risky. We study the risk facing a well-documented and important class of entrepreneurs, those backed by venture capital. Using a dynamic program, we calculate the certainty-equivalent of the difference between the cash rewards that entrepreneurs actually received over the past 20 years and the cash that entrepreneurs would have received from a risk-free salaried job. The payoff to a venture-backed entrepreneur comprises a below-market salary and a share of the equity value of the company when it goes public or is acquired. We find that the typical venture-backed entrepreneur received an average of $5.8 million in exit cash. Almost three-quarters of entrepreneurs receive nothing at exit and a few receive over a billion dollars. Because of the extreme dispersion of payoffs, an entrepreneur with a coefficient of relative risk aversion of two places a certainty equivalent value only slightly greater than zero on the distribution of outcomes she faces at the time of her company’s launch. (JEL G24, G32, L26, M13)

An entrepreneur’s primary incentive is ownership of a substantial share of the enterprise that commercializes the entrepreneur’s ideas. An inescapable consequence of this incentive is the entrepreneur’s exposure to the idiosyncratic risk of the enterprise. Diversification or insurance to ameliorate the risk would necessarily weaken the incentives for success.

We study this issue in the case of startup companies backed by venture capital. These startups are mainly in information technology and biotechnology. They harness teams comprising entrepreneurs (scientists, engineers, and executives), venture capitalists (general partners of venture funds), and the suppliers of capital (the limited partners of venture funds). During the startup process, entrepreneurs collect only submarket salaries. The compensation that attracts them to startups is the share they receive of the value of a company if it goes public or is acquired.

We make use of a rich body of data, which covers close to the universe of companies receiving venture funding from 1987 to 2008, though some information is missing for many companies. We use a method for imputing missing data that takes account of selection bias.

Our most important finding is that the reward to the entrepreneurs who provide the ideas and long hours of hard work in these startups is zero in almost three quarters of the outcomes, and small on average once idiosyncratic risk is taken into consideration.

Standard venture deals involve three parties—entrepreneurs, general partners, and limited partners. The entrepreneurs have leveraged positions; that is, they receive no payoff until other claimants have received prescribed payoffs. The general partners, who arrange financing and supervise the startup company by holding board seats, are compensated in proportion to the amount invested and the capital gains from the investment. The limited partners are passive.

* Hall: Hoover Institution and Department of Economics, Stanford University, 434 Galvez Mall, Stanford, CA 94305-6072 (e-mail: rehall@stanford.edu); Woodward: Sand Hill Econometrics, 115 Everett Street, Palo Alto, CA 94301 (e-mail: swoodward@sandhillcon.com). We are grateful to the referees, Ravi Jagannathan, Deborah Lucas, Matthew Rhodes-Kropf, and numerous seminar participants for comments and to Katherine Litvak for data on venture contract terms. Hall’s research is part of the Economic Fluctuations and Growth Program of the National Bureau of Economic Research.
investors who hold debt and equity claims on the startup. General partners are somewhat diversified across investments, and the limited partners are highly diversified. The burden of specialization falls mainly on the entrepreneurs. Robert E. Hall and Susan E. Woodward (2007) describe the returns to the general partners and limited partners. This paper deals exclusively with the entrepreneurs.

Although the average ultimate cash reward to an entrepreneur in a company that succeeds in landing venture funding is $5.8 million, most of this expected value comes from the small probability of a great success. An individual with a coefficient of relative risk aversion of two and assets of $188,949 is indifferent between employment at a market salary and entrepreneurship. With lower risk aversion or higher initial assets, the entrepreneurial opportunity is worth more than alternative employment. We infer that entrepreneurs are drawn differentially from individuals with lower risk aversion and higher assets. Other types of people that may be attracted to entrepreneurship are those with preferences for that role over employment and those who exaggerate the likely payoffs of their own products. Our model does not include these factors, however—we use standard preferences based on consumption levels alone.

We focus on the joint distribution of the duration of the entrepreneur’s involvement in a startup—what we call the venture lifetime—and the value that the entrepreneur receives when the company exits the venture portfolio. Exits take three forms: (i) an initial public offering, in which the entrepreneur receives liquid publicly traded shares six months after the IPO and has the opportunity to diversify; (ii) the sale of the company to an acquirer, in which the entrepreneur receives cash or publicly traded shares in the acquiring company and has the opportunity to diversify; and (iii) shutdown or other determination that the entrepreneur’s equity interest has essentially no value. Most IPOs return substantial value to an entrepreneur. Some acquisitions also return substantial value, while others may deliver a meager or zero value.

The joint distribution shows a distinct negative correlation between exit value and venture lifetime. Highly successful products tend to result in IPOs or acquisitions at high values relatively quickly. These outcomes are favorable for entrepreneurs in two ways. First, the value arrives quickly and is subject to less discounting. Second, the entrepreneur spends less time being paid a low startup salary and correspondingly more time with higher poststartup compensation, in the public version of the original company, in the acquiring company, or in another job. A fraction of entrepreneurs launch new startups after exiting from an earlier startup.

Throughout the paper, we study exit values from the point of view of the individual entrepreneur. About a quarter of entrepreneurs do not share the proceeds with other entrepreneurs; they operate solo. Another quarter share the entrepreneurial role equally with another founder. In the remaining cases, entrepreneurial ownership is distributed asymmetrically between a pair of entrepreneurs, or there are three or even more entrepreneurs. We measure the total entrepreneurial return delivered by each venture company and then infer the returns to the individual entrepreneurs from information about the distribution among entrepreneurs. All of the tabulations in the paper refer to entrepreneurs, not to companies.

We develop a unified analysis of the factors affecting the entrepreneur’s risk-adjusted payoff, based on a dynamic program. The analysis takes account of the joint distribution of exit value and venture lifetime and of salary and compensation income. We use it to calculate the certainty-equivalent value of the entrepreneurial opportunity—the amount that a prospective entrepreneur would be willing to pay to become a founder of a venture-backed startup. For a risk-neutral individual, the certainty equivalent is $5.8 million. With mild risk aversion and savings of $100,000, however, the amount is only $0.6 million, and with normal risk aversion and that amount of savings, the certainty equivalent is slightly negative.

We are not aware of any earlier research that quantifies the rewards on a per-company basis or per-entrepreneur-year basis, the focus of our work. Earlier research on venture backed startups
has focused on the returns to venture investors. An extensive theoretical literature considers
the implications of idiosyncratic risk for entrepreneurs and managers—see John Heaton and
Deborah Lucas (2004) for a recent contribution and many references.

I. What This Paper Does

Our first step is the development of data for the great majority of all venture-backed start-
ups in the United States from the date of each one’s first venture funding, covering subsequent
venture fundings, the date of exit, and the cash payoff to the entrepreneurs as a group. We start
by determining the founding equity interest of the entrepreneurs, then track the dilution of the
entrepreneurs’ interest through successive rounds of venture investment. We take account of
extra dilution mandated by the standard venture contract if a later round of investment places
a lower value on the company than did an earlier round—a “down round.” We calculate the
entrepreneurs’ share of the proceeds from an exit. Here we take account of the debt aspects of
the venture investors’ claims on the company, the preferences associated with their holdings.
Successful exits are IPOs or high-priced acquisitions by another company; unsuccessful ones
are shutdowns and low price acquisitions in which the entrepreneurs receive nothing. The output
from the first step is the starting date, exit date, and combined exit value for the entrepreneurs in
each of the companies.

The next step develops tabulations of the joint distribution of venture lifetime—the time from
startup to exit—and exit values for individual entrepreneurs, based on data for entrepreneurs’
group exit values and data on the distribution of individual entrepreneurs’ shares among the total
equity holdings of the entrepreneurs as a group.

The third step constructs a personal financial model of the entrepreneur during the startup
experience. The model captures the implications of the idiosyncratic risk facing the entrepreneur.
No insurance is available to deal with the risk. Rather, the entrepreneur works for an uncertain
number of years at a submarket salary and may or may not receive an exit value somewhere
between a few hundred thousand dollars and a billion dollars. The model considers both ele-
ments of risk—the number of years of foregone earnings and the uncertainty about the payoff.

II. The Startup Process

At the outset, startups are usually operated and financed by the entrepreneurs themselves.
Friends and family may invest as founding shareholders. Unless the founders are wealthy, they
need outside financing, so a main task early in a startup is to find investors. Some are individual
investors called angels. But venture funds are capable of investing more at the outset than is
available from these other sources, and venture can invest large amounts later in the development
of a startup with a promising product. Our concern is with the companies that succeed in obtain-
ing venture funding by convincing some venture capitalists that the new business has a positive
net present value, which, given the skewness of the distribution of value at outcome, implies at
least some chance of becoming highly profitable.

Venture funds seldom give a company all of the money it will need to get from startup to exit
in a single investment. Instead, a syndicate of venture funds will provide financing in rounds,
anticipating future rounds of funding, possibly including different investors, if the startup makes
reasonable progress but still lacks the revenue to be self-sustaining, and denying the startup fur-
ther funding otherwise. An early round typically gives a startup a few million dollars, while later
rounds, if they occur, often involve much larger investments.

General partners are the organizers of venture funds. They recruit financing commitments from
limited partners—usually pension funds, endowments, and wealthy individuals—and choose the
companies that will receive financing. Compensation to general partners comprises an annual fraction of two to three percent of the limited partners’ invested capital plus carry—20 to 30 percent of the profit from successful exits. The limited partners receive most of the cash returned by venture investments when a company undergoes a favorable exit event—an IPO or acquisition.

Venture funds generally hold convertible preferred shares in their portfolio companies. The preference requires that the funds receive a specified amount of cash back before the common shareholders (the entrepreneurs, angels, and employees) receive any return. In a successful outcome, the convertible preferred shares convert shares to common stock. Instead of convertible preferences, venture funds may hold debt claims, in which case they receive the repayment of the debt even in the best outcomes. Both arrangements put the common shareholders, including the entrepreneurs, in a leveraged position, increasing their exposure to the idiosyncratic risk of the startup.

A huge literature portrays the standard venture financial contract as the constrained optimum of a challenging mechanism design problem. This research explains key features, including the assignment of a share of the ultimate value to the entrepreneurs, multiple stages of financing, and debt instruments (preferences) that convert to equity. Some of the more prominent contributions include Anat R. Admati and Paul Pfleiderer (1994); Klaus M. Schmidt (2003); Catherine Casamatta (2003); and Rafael Reullo and Javier Suarez (2004). Alex Wilmerding (2003) and Constance E. Bagley and Craig E. Dauchy (2003) explain the terms of venture contracts from the perspective of venture capitalists and their lawyers.

The dominant factor in this literature is moral hazard. Venture investors and their agents, the general partners of venture funds, are unable to monitor or specify the efforts of entrepreneurs to commercialize their ideas. Consequently, the entrepreneurs are paid in proportion to the actual commercial success of their companies. This alignment of incentives comes at the cost of a substantial diminution in the value of the enterprise because of the idiosyncratic risk that entrepreneurs are unable to insure. Alternatives with less risk, such as paying entrepreneurs salaries in place of equity, apparently provide such weak incentives that the relationship based on equity incentives weakened by idiosyncratic risk is still optimal for some products and some entrepreneurs.

Venture capitalists face a daunting problem evaluating proposals for startups. One of the reasons that entrepreneurs receive submarket salaries during the startup phase is to induce self-selection among applicants for venture funding. Only entrepreneurs with confidence in the commercial values of their ideas will seek funding if the entrepreneur’s payoff from an unsuccessful startup is negative.

Most of the expected return to entrepreneurs comes from low-probability large gains. About three-quarters of venture-backed companies expire without returning any cash to their entrepreneurs. The largest returns generally come from IPOs, but acquisitions sometimes provide high returns as well. On the other hand, many acquisitions occur at low prices and are effectively liquidations. Some venture-backed companies remain for many years as stand-alone operations, able to pay their employees out of revenue but generating no returns for shareholders.

The free-standing startup company is one of the ways that ideas for new products are developed and marketed. It provides powerful incentives for its entrepreneurs, but at the cost of exposing them to the idiosyncratic risks of their companies. Most scientists and engineers working on new products work as employees for established—often very large—companies. Their employment contracts isolate them from most of the idiosyncratic risks of the products they develop. Incentives are not as powerful as in startups. We discuss the sorting of potential entrepreneurs into startups and established companies in a concluding section. We note that the market for scientists and engineers has not developed any intermediate contract, though one could imagine such a contract. It would pay a higher salary than the standard venture contract does but provide
less exit value, for example, by putting a ceiling on the payout. We believe that such contracts are rare. The two successful contract forms in the market for technical talent are polar opposites. The intermediate contract appears not to be viable.

III. Data

A. Data on Venture Transactions

We use a database compiled by Sand Hill Econometrics on venture investments in startups and on the fates of venture-backed companies. The data are drawn from a variety of sources, including several commercial data vendors. A comparison of the list of companies in the database to lists of companies in pension plan venture investments shows that the Sand Hill data includes close to the universe of venture-funded startups from 1987 to the present. The data vendors concentrate on reporting funding events and valuations for venture investments and successful outcomes (IPOs and high value acquisitions) and are less likely to report shutdowns and acquisitions at low values. Sand Hill Econometrics has used a wide range of sources to augment coverage of these adverse termination events. The Data Appendix posted in the archives of this journal and on the first author’s Web site describes the data in more detail and documents the technique we use to track the evolution of the entrepreneurs’ ownership of a company through successive rounds of funding, each of which dilutes the entrepreneurs’ claims.

Our measurement of the entrepreneur’s take from the exit value of the enterprise starts with the total cash received by the owners collectively. In the case of an IPO, this amount is the total market value of the newly public company less the cash raised in the IPO. For an acquisition, it is the total amount paid to the shareholders of the company. We divide this amount among the owners in the way specified in the standard venture contracts between the venture capitalist and shareholders.

Immediately prior to the first venture investment, the entrepreneurs own most of the enterprise. The other shareholders are usually angels—individual investors—and friends and family. We assume that the cash investments from the entrepreneurs are made prior to the first round of venture investments. As the development of the company progresses, the entrepreneurs’ ownership share declines. The main reason is that each round of venture investment purchases equity and debt claims that dilute the entrepreneur. In addition, the typical startup hires professional managers who receive stock options that convert to share ownership upon a remunerative exit event.

For initial shares owned by angels, friends-and-family, and executives, we use data from published studies that report averages across venture-backed startups. For dilution from venture investments, we use data specific to each company based on round by round data from the Sand Hill database.

Each round of venture financing purchases equity and debt claims on the startup. The debt claims take the form of preferences—cash due to venture investors upon an exit event on top of their equity claims. Some preferences pay off only in the poorer outcomes, while others pay off in all outcomes. We apply standard formulas from venture contracts to estimate the deductions from entrepreneurial receipts resulting from preferences. The new equity issued in a round dilutes the ownership shares of the entrepreneurs. For investment rounds where the purchase price of the new shares—and thus the current value of all shares—are reported in the database, the dilution calculation is straightforward. Where the purchase price is not reported, we estimate the share of the company’s equity purchased in the round by the investors using a body of data suited to solving the sample selection problem.
Another feature of the standard venture contract is antidilution protection to venture investors from earlier rounds if a later round assigns a lower value to the company. This protection shifts ownership from the entrepreneurs to the earlier venture investors to eliminate or ameliorate the decline in value they would otherwise suffer from a so-called down round.

One important source of valuation data is S-1 statements filed by venture-backed companies when they go public. These statements often give a funding history for the company. Because an IPO is a favorable event, the back-filling of round values from S-1s is a source of return-based selection in the data. The Appendix describes how we adjust for selection bias.

Our data include 22,004 venture-backed companies, the great majority of all such companies in the United States for the period from the beginning of 1987 through the third quarter of 2008. Among the exit values used in the analysis, 2,015 are IPOs, 5,625 are acquisitions, and 3,352 are confirmed zero value exits. Of the remaining companies, we treat those more than five years past their last rounds of venture funding as having exited at some time with zero value; 4,220 companies fall into this category. We randomly assign these companies exit dates by drawing from the empirical distribution of time past funding of companies with known zero value exit dates. The remaining 6,792 companies have not yet achieved their exit values.

For acquisitions, we use the reported exit value and exit date as the entrepreneur’s payoff, as we believe that lags in payments to entrepreneurs are quite brief. For IPOs, we assume that entrepreneurs are required to retain all of their publicly traded shares for a lockup period of six months, so we date the receipt of cash at six months past the IPO and the amount as the market value of the entrepreneurs’ holdings at that date.

We state all exit values in 2006 dollars using the Consumer Price Index.

Finally, we use the National Bureau of Economic Research (NBER) TAXSIM model (http://www.nber.org/~taxsim) to calculate the after-tax value of the cash received by an entrepreneur by applying the marginal tax rate on long-term capital gains to the entrepreneur’s exit cash, under US and California income taxation (the majority of the entrepreneurs in the sample live in California). The rate is very close to 25 percent at all relevant levels of salary and capital gains income. We use 25 percent in all cases. We also use TAXSIM to calculate the after tax values of the venture and alternative salaries. We consider pretax salaries of $150,000, $300,000, $600,000, and $2,000,000, which correspond to $111,220, $194,126, $367,212, and $1,128,001, after tax. We use 2006 tax rates, which are essentially the same as the rates for other recent years.

B. Share of Ownership by Individual Entrepreneur

We use a model of personal or family decision making, where consumption depends on the earnings and exit values of individuals. Our data treat all the entrepreneurs in a company as a group. Our basic data sources do not contain information about the ownership shares of the individual entrepreneurs in each startup company. We use estimates from a sample of companies that underwent IPOs. The sample is a random draw of 100 candidates from all IPOs reported in our data. The SEC form S-1 filed prior to an IPO often contains a description of the major shareholders, which includes the entrepreneurs. The sample contains 41 companies and 66 entrepreneurs. Because the venture IPO sample is not necessarily representative of venture-backed startups in general, we regard it as illustrative and far from definitive. We see little benefit from extending the sample to a larger number of IPOs. Our results are not at all sensitive to the method for restating company based distributions as entrepreneur based distributions.

Table 1 describes the venture IPO sample. Just under a quarter of entrepreneurs receive all of the entrepreneurial exit value—these are solo entrepreneurs. At the other end, about a sixth of entrepreneurs receive less than 20 percent of the exit values of their companies. The
right-hand column of the table shows the exit value of all of the entrepreneurs averaged across all companies that contain an entrepreneur in the share category corresponding to the row in the table. A solo entrepreneur, in the bottom line of the table, receives all of an average of $91 million of exit value, while an entrepreneur with less than a 20-percent share receives less than a fifth of an average of $48 million of exit value. There appears to be a positive relation between an entrepreneur’s share of the entrepreneurial exit value and the magnitude of that value, within the IPO sample.

Table 1 suggests that, among IPO exits, a solo entrepreneur is likely to be affiliated with a company with a somewhat higher exit value than other entrepreneurs. In our framework, we encounter this issue the other way around—we need the distribution of entrepreneurial shares conditional on the size of the exit. The distribution of individual entrepreneur’s exit value depends on the joint distribution of the two variables. The individual’s exit cash is the product of individual’s share and the total exit value.

We consider two cases. Our base case assumes independence of the total entrepreneurial exit value and the share of that value received by a particular entrepreneur. Our alternative case emulates the joint relationship shown in [Table 1]. In both cases, we constrain the marginal distribution of the entrepreneur’s share to be the distribution shown in [Table 1]. We use the empirical distribution of total entrepreneurial exit value derived from the database. Because we impose the same prescribed marginal distributions of the two variables, our two cases differ only in the copula of the joint distribution. The Appendix describes our procedure for finding the joint distribution for the alternative case. For our base case with independence, the joint distribution is simply the product of the marginals.

IV. The Joint Distribution of Startup Lifetime and Exit Value

The lifetime of a startup—the time from inception to the entrepreneurs’ receipt of cash from an exit event—plays a key role in our analysis. Entrepreneurs prefer short lifetimes for two reasons. First, their salaries at a venture-backed startup are modest; they forgo a full return to their human capital during the lifetime. Second, the time value of money places a higher value on cash received sooner.

Lifetimes and exit values are not distributed independently. In particular, a substantial fraction of startups linger for many years and then never deliver much cash to their founders. And some of the highest exit values occurred for companies like YouTube that exited soon after inception.

Our calculations also need to make the transition from data based on companies to distributions over entrepreneurs, as discussed above. We start with the joint cumulative distribution, \( F_{\tau}(v_c) \), of startup lifetime, \( \tau \), and value received by the company, \( v_c \). We have the discrete distribution,
The cumulative distribution of the entrepreneur’s exit value, $v$, is

$$G_v(v) = \sum_i F_\tau \left( \frac{v}{s_i} \right) h_i.$$  

Here $s_i$ is the average entrepreneurial fraction in category $i$, shown in the fourth column of Table 1. In words, the joint probability for a range of values of the entrepreneur’s exit value, $v$, say from $v'$ to $v''$, and venture lifetime, $\tau$, is the sum over the distribution of entrepreneur’s shares, $h_i$, of the fraction of company exit values in the range from $v'/s_i$ to $v''/s_i$. Another way to express the range is that the company exit value multiplied by the share, $v_c s_i$, lies in range from $v'$ to $v''$.

We take a flexible view of the joint distribution, as appropriate for our rich body of data. We place lifetimes $\tau$ and values $v$ in nine and 11 bins respectively and estimate the 99 values of the joint distribution defined over the bins. Estimation of the joint distribution needs to take account of the fact that many companies in our data have not completed their lifetimes as startups. To account for the right-censoring of lifetimes, we let $I_{c,\tau}$ be an indicator function for whether a company started in month $t$ could have been observed to exit at lifetime $\tau$. We denote the month where we gather our data as $T$. Thus

$$I_{c,\tau} = \begin{cases} 1 & \text{if } T - t \geq \tau \\ 0 & \text{otherwise.} \end{cases}$$

We further let $N_{v,\tau}$ be the number of entrepreneurs in the sample with entrepreneurial exit value in bin $v$ and lifetime in bin $\tau$. That is,

$$N_{v,\tau} = \sum_i \text{Count}(v, i) M h_i$$

where $\text{Count}(v, i)$ is the number of companies whose exit value $v_c$ is such that $v_c s_i$ falls in entrepreneurial exit value bin $v$ and $M$ is the average number of entrepreneurs per company. Entrepreneurs from nonexited companies are not included in $N$. We let $L_t$ be the number of companies launched in month $t$. Then

$$N_{v,\tau} = \sum_t M L_t I_{t,\tau} g_{v,\tau},$$

where $g$ is the discrete joint distribution defined over the bins, the differences in the cumulative distribution $G$. We let

$$\tilde{N}_{\tau} = \sum_t M L_t I_{t,\tau},$$

so

$$g_{v,\tau} = \frac{N_{v,\tau}}{\tilde{N}_{\tau}}.$$ 

Our method for estimating the joint distribution is equivalent to estimating a hazard function showing the probability of exit at a given age conditional on no earlier exit, using all available data on the hazard at each age.

This approach to estimating the joint distribution does not constrain it to sum to one. In our data, the sum is 0.83. Any reasonable approach to imposing that constraint could be rationalized...
as the minimization of some weighted distance function. We choose the obvious one, which is to divide the distribution from (6) by the sum of all of its values. Figure 1 shows the estimated joint distribution. The left row, with literally zero exit value to the entrepreneur, dominates the probability. Most of the remaining probability goes to moderate exit values with relatively brief lifetimes. Exit values above $100 million are quite rare. Table 2 shows the joint distribution numerically, along with the marginal distributions of exit value and venture lifetime. We find 21 instances where the entrepreneur received at least $100 million and the venture lifetime was 12 months or less—of these, 9 were IPOs and the remaining 12 acquisitions.

The marginal distributions shown in Figure 2 and Figure 3 and tabulated in Table 2 provide useful alternative views of the joint distribution. Figure 2 shows the marginal distribution of exit value, summed across all the lifetime categories. It shows that 75 percent of all startups deliver zero exit value. Categories of low but positive exit value account for most of the rest of the outcomes. Only a tiny fraction of entrepreneurs receive more than $100 million in exit value. Figure 3 shows the conditional distribution of lifetime given exit value. Each row sums to one. Note that the two axes on the floor are reversed relative to Figure 2 to make it easier to see the shape of the distribution. This figure shows the negative correlation of lifetime and exit value. At the front, the figure shows that zero-value exits tend to have long lifetimes. At the back, it shows that high-value exits tend to have short lifetimes. The conditional distributions of the high-value exits are irregular because there are few of them, though they account for a significant fraction of the total exit value.

The fraction of entrepreneurs who received nothing in equity value from their efforts was large throughout the period covered by our data, ranging from 58 percent in 2006 through 2008 to 87 percent in 1999 through 2001. The conventional wisdom that only about half of entrepreneurs fail to receive any equity return is inconsistent with our findings.
### Table 2—Joint Distribution of Venture Lifetime and Exit Value, Percent Probability by Cell

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<th>Exit value (millions of dollars)</th>
<th>Venture lifetime, years</th>
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<td>All</td>
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<table>
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<th>7 to 9</th>
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![Figure 2. Marginal Distribution of Exit Value](image-url)
Figure 3. Conditional Distribution of Venture Lifetime Given Exit Value

Figure 4. Fractions of Total Exit Value by Exit Value Category
Figure 4 shows the distribution of the total exit value facing an entrepreneur, by exit value category. Each bar shows the fraction of the total arising within the category. The category contributing the greatest fraction of value is a billion dollars or more, despite the rarity of such payoffs. Figure 5 shows the marginal distribution of venture lifetimes. The modal lifetime is between one and three years. The median is somewhat above four years. We do not calculate a mean lifetime, because the mean is sensitive to the extreme values, which are difficult to measure. Figure 6 shows the distribution of exit value by lifetime. More than a quarter of the total value arises from companies with venture lifetimes between one and two years. Not only is this range of lifetimes common, but exits that soon tend to have higher values.

V. Economic Payoffs to Entrepreneurs

Venture-backed companies typically have a scientist or similar expert, or a small group, who supply the original concept, contribute a small amount of capital, and find investors to supply the bulk of the capital. These entrepreneurs, together with any angels, own all of the shares in the company prior to the first round of venture funding.

The entrepreneurs are specialized in ownership of the venture stage firm. Our approach to valuation takes account of the heavy exposure of the entrepreneur to the idiosyncratic volatility of the company. We also take account of the modest salaries that entrepreneurs generally receive during the venture phase of the development of their companies and of the lifetime of the company, which affects the discounting applied to the exit value and the burden of the low salary.

Our model assumes that the entrepreneurs in a company have already made all of their financial investments in their company; all further funds will come from venture investors. We believe this assumption to be generally realistic, though of course some entrepreneurs are able to continue financing their companies alongside venture investors. We portray an entrepreneur as having some savings available to finance consumption beyond what the relatively low venture salary will support. We rule out the possibility that an entrepreneur could borrow against future earnings or against the possible exit value of the company. We are quite sure this assumption is
realistic. Thus the entrepreneur makes a decision each year about how much to draw down savings during the year; that is, by how much consumption will exceed the venture salary.

A. Analytical Framework

Our framework starts from a standard specification of intertemporal preferences for entrepreneurs—they order random consumption paths according to

\[ \mathbb{E} \sum_t \left( \frac{1}{1 + r} \right)^r u(c_t). \]

Here \( r \) is the entrepreneur’s rate of time preference and the rate of return on assets; \( u(c) \) is a concave period utility. We define the function \( U(W) \) as the utility from a constant path of consumption funded by wealth \( W \):

\[ U(W) = \frac{1 + r}{r} u \left( \frac{r}{1 + r} W \right). \]

The multiplication by \( (1 + r)/r \) turns flow utility into discounted lifetime utility. The quantity \( r/(1 + r) W \) is the flow of consumption to be financed by the return on the wealth at rate \( r \).

We distinguish between wealth, \( W_t \), which measures the entrepreneur’s total command over resources, and so incorporates the expected value of future compensation (human wealth), and assets, \( A_t \), by which we mean holdings of nonhuman wealth as savings. \( A_t \) does not include the entrepreneur’s holdings of shares in the startup, which we classify as human capital. For an entrepreneur in year \( t \) of a startup that has not yet exited, we define \( W_t(A_t) \) as the wealth equivalent of the entrepreneur’s command over resources, counting what remains of the entrepreneur’s original nonhuman wealth, \( A_t \), and the entrepreneur’s random future payoff from the startup,
conditional on not having exited to this time. Our definition is implicit: $U(W_t(A_t))$ is the expected utility from maximizing equation (7) over consumption strategies.

Now we let $U(W_t(A_t))$ be the value, in utility units, associated with an entrepreneur in a non-exited company $t$ years past venture funding, as a function of current nonentrepreneurial assets $A_t$. We could have defined a value function $U_t(A_t)$ without interposing the function $W_t(A_t)$. Instead we let $W_t(A_t)$ be the value function, which means that we need to take the concave transformation $U(W_t(A_t))$ so that the Bellman equation adds up utility, according to the principle of expected utility. The slightly roundabout approach of stating our findings in terms of the wealth equivalent $W_t(A_t)$ makes the units meaningful, whereas the units of utility are not. Further, in our benchmark case, utility is negative, a further source of confusion. Note that $W$ captures initial assets, venture salary, venture exit value, and subsequent compensation in a postventure position, when it is calculated at time zero for an entrepreneur.

The company has a conditional probability or hazard $\pi_t$ of exiting at age $t$. At exit, it pays a random amount $X_t$ to the entrepreneur. Upon exiting, the entrepreneur’s value function is $U(W^*_t(A_t))$, where $A$ now includes the cash exit value. The entrepreneur’s consumption is limited by assets left from the previous year—no borrowing against future earnings may occur. The entrepreneur’s dynamic program is

$$
U(W_t(A_t)) = \max_{c_t < A_t} \left[ u(c_t) + \frac{1}{1 + r} \left( \pi_{t+1} U(W_{t+1}((A_t - c_t)(1 + r) + w)) + \frac{1}{1 + r} \pi_{t+1} \mathbb{E}_X U(W^*((A_t - c_t)(1 + r) + X_{t+1})) \right) \right].
$$

The postventure value function is

$$
U(W^*(A)) = \frac{1 + r}{r} u \left( \frac{rA + w^*}{1 + r} \right).
$$

Here $w^*$ is postventure compensation, including employee stock options, at the nonventure continuation of this company or another company. From equations (8) and (10), we have

$$
W^*(A) = A + \frac{w^*}{r}.
$$

Note that this is additive in $A$. But when future earnings are random, the entrepreneur’s risk aversion enters the calculation of the wealth equivalent.

We represent each of the value functions $U(W_t(A_t))$ as piecewise linear with 500 knots between zero and $49$ million, spaced exponentially. We calculate them by backward recursion (value function iteration). We assume power utility with constant relative risk aversion, $\gamma$. We take as our base case $\gamma = 2$, a venture salary $w$ equal to the posttax value of $150,000, postventure compensation $w^*$ equal to the posttax value of $300,000, and starting assets of $A_0 = 1$ million.

A useful feature of the wealth equivalent is that the difference between its value for an entrepreneur with given initial assets and its value for an individual who holds a nonventure position paying $w^*$ and with the same initial assets is the amount that the second would be willing to pay to become an entrepreneur. We call this the certainty equivalent value of the entrepreneurial opportunity and denote it $\tilde{A}$. This property follows from the additivity of the nonentrepreneurial wealth equivalent we noted earlier.
B. Results

Figure 7 shows $W_0(A_0)$, the wealth equivalent for an entrepreneurial experience as of its beginning and $W^*(A_0)$, the wealth equivalent for a nonentrepreneur, both as functions of the common value of their initial assets, shown on the horizontal axis. The certainty equivalent value of the venture opportunity is the vertical difference between the two curves. The nonentrepreneurial value is a straight line with unit slope—a dollar of extra initial assets becomes a dollar of wealth, because we assume that the nonventure individual faces no uncertainty. On the other hand, a dollar of extra initial assets becomes more than a dollar of equivalent wealth, because initial wealth has no uncertainty and thus dilutes the uncertainty from the venture outcome. This property is a cousin of the principle that people should treat risky outcomes as if they were worth essentially their expected values, when the outcomes are tiny in relation to their wealth. The slope of the entrepreneur’s value is more than three at low levels of assets but declines to 1.03 at assets of $20$ million.

The figure shows that, despite the chance of making hundreds of millions of dollars in a startup, the economic advantage of entrepreneurship over an alternative career is not substantial. The burden of the idiosyncratic risk of a startup falls most heavily on those with low initial assets. The entrepreneur with less than a million dollars of initial assets faces a heavy burden from the risk and has a lower certainty equivalent wealth than the nonentrepreneur.

Table 3 gives the certainty equivalent value of the entrepreneurial opportunity for 36 combinations of the three determinants: the coefficient of relative risk aversion, the compensation at an alternative, nonentrepreneurial job, and the entrepreneur’s assets at the beginning of entrepreneurship. The first three lines take the entrepreneur to be risk neutral, so the values are just present values at the five-percent annual real discount rate. In this case, the value is the same for any level of initial assets. The value is $5.8$ million. The value is $5.1$ million for an individual with a nonentrepreneurial opportunity to earn $600,000 per year before tax. If the nonentrepreneurial opportunity pays $2$ million per year before tax, the venture opportunity has barely positive value. A typical startup probably cannot attract an established top executive from a large public corporation, even if the executive is risk neutral, as his earnings are generally even higher than $2$ million.
Table A4 in the Appendix gives estimated standard errors of the figures in Table 3. They are sufficiently small that none of our conclusions is much clouded by sampling variation.

The conclusions from the table are similar if the individual is mildly risk averse, with a coefficient of relative risk aversion of 0.9. The advantage of the entrepreneurial opportunity, stated as a wealth equivalent, is only $0.6 million for an entrepreneur with $0.1 million in assets and only $1.2 million for an entrepreneur with $5 million. These figures are negative or only slightly positive if the nonentrepreneurial opportunity pays $600,000 per year before tax.

At the standard value of the coefficient of relative risk aversion, 2, the advantage of the entrepreneurial opportunity is generally small or negative—deeply negative if the nonentrepreneurial opportunity pays $2 million per year. In our base case, with nonentrepreneurial compensation of $300,000 per year before tax and $1 million in assets, the advantage of the entrepreneurial opportunity is only $0.2 million. The incentive is not impressive for larger asset holdings. With higher compensation at the nonentrepreneurial job, the advantage disappears unless the individual is quite rich.

C. Implications of Correlation between the Number of Entrepreneurs in a Company and its Exit Value

The results earlier in this section rest on the assumption that companies with more entrepreneurs have the same distribution of exit values as do other companies. We have solved the dynamic program in our base case with an alternative specification that implies a ratio greater than one of average exit value of companies affiliated with a single entrepreneur to the average exit value of companies affiliated with an entrepreneur with less than a 20 percent share of total entrepreneurial equity. The third column of Table 1 shows that the ratio is 1.9 within our sample of IPOs. Our alternative specification matches that ratio.

The solution to the dynamic program is very similar with the alternative specification. In our base case, the value of initial savings $A_0$ that makes the entrepreneur indifferent between entrepreneurship and the outside salary opportunity is $188,949, while the indifference point for the alternative is $184,352.

VI. Entrepreneurs in Aging Companies

Our discussion so far has focused on the risk adjusted payoff to a potential entrepreneur at the decision point when venture funding first becomes available. In this section, we consider

<table>
<thead>
<tr>
<th>Coefficient of relative risk aversion, $\gamma$</th>
<th>Pretax compensation at nonentrepreneurial job, thousands of dollars per year</th>
<th>Certainty equivalent of entrepreneurial opportunity, millions of dollars</th>
<th>Assets at beginning, millions of dollars</th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>300</td>
<td>5.8</td>
<td>5.8</td>
</tr>
<tr>
<td>0</td>
<td>600</td>
<td>5.1</td>
<td>5.1</td>
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<tr>
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<td>2.3</td>
<td>2.3</td>
</tr>
<tr>
<td>0.9</td>
<td>300</td>
<td>0.6</td>
<td>1.2</td>
</tr>
<tr>
<td>0.9</td>
<td>600</td>
<td>-0.2</td>
<td>0.3</td>
</tr>
<tr>
<td>0.9</td>
<td>2,000</td>
<td>-5.8</td>
<td>-4.0</td>
</tr>
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<td>600</td>
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<tr>
<td>2</td>
<td>2,000</td>
<td>-13.9</td>
<td>-8.9</td>
</tr>
</tbody>
</table>

Table 3—Certainty Equivalent Value of the Venture Opportunity
the same issue at later decision points, as the startup ages. Our discussion is conditional on the company not having exited.

The dynamic program of equation (9) assigns a value $W_t(A_t)$ to the entrepreneur’s position in each year $t$ that the company has not exited. Under our assumptions, the path is the same for all companies. The entrepreneur’s value falls as the company ages for two reasons. First, the entrepreneur generally consumes out of assets, so assets decline. Second, early exits are the most valuable exits, so aging another year means that the remaining potential exit values are not as valuable. Figure 8 shows the path of $W_t(A_t)$. It declines from $5.1$ million at the outset to $4.3$ million at age 10, conditional on no exit. From that point, the value rises, because the distribution of exit values becomes more favorable, though not as favorable as for young startups.

The figure also shows the individual’s value of a nonentrepreneurial job, $W^*(A_t)$. It declines as well, but only for the first reason, the drawdown of assets to finance consumption in excess of the low startup salary.

Figure 9 shows the paths of assets and consumption as a company ages. For the first decade, assets decline because consumption exceeds the modest startup salary and the entrepreneur has no other source of current cash, pending a favorable exit. During this period consumption declines, because, as an exit fails to occur during the early years, the entrepreneur learns that risk-adjusted well-being, as measured by $W_t(A_t)$, has declined. Eventually assets fall to the point of consumption. From this point until exit, the entrepreneur lives on the salary and maintains assets only as a way to spread consumption between paychecks (we assume, for simplicity, that the entrepreneur receives the salary at the end of each year, and we measure assets at the beginning of the year). The line labeled $c(W_t)$ shows the level of consumption that a consumer without a cash flow constraint would choose, given lifetime prospects as measured by $W_t$. Consumption starts out only slightly below this level, but as the entrepreneur depletes assets, consumption falls toward the cash flow limit. In the event that the startup ages into its second decade, the cash flow constraint keeps consumption far below its unconstrained level.
VII. Robustness and Related Results

Table 4 shows certainty equivalent values for a variety of alternative specifications. The first line repeats the base case from Table 3. The second line limits the companies included in the calculations to those in information technology. In all but the top initial wealth category, the values are slightly lower for these companies than for venture-backed companies in general. For the less risk-averse entrepreneurs in the top wealth category, the value is higher, reflecting the disproportionate role of IT in the most successful startups, such as Google. The third line limits the companies to those in biotech. The certainty equivalent value of a biotech startup is higher in every wealth category.

The fourth and fifth lines of Table 4 break the sample by the date of first venture funding. The basic conclusions of the paper apply equally to companies launched in 1995 and earlier and to those launched after 1995.

Line 6 of the table explores the sensitivity of our calculations to our procedure for imputing the dilution of the entrepreneur’s share in a round of venture funding when the dilution is not reported directly. We add one percent to the fitted value from the logit regression that imputes the share of the company in all such rounds sold to that round’s group of venture investors. The extra dilution depresses the entrepreneur’s certainty equivalent value, but the change does not affect any of our conclusions.

Line 7 calculates the effect of the provisions of the standard venture contract that shift ownership to the investors at the expense of the entrepreneurs and other common shareholders in the case of a down round. In such a round, the share price is below the price of an earlier round. The increase is tiny—the great bulk of payoffs to entrepreneurs come from IPOs and higher value acquisitions, and companies enjoying these favorable outcomes rarely experience down rounds.
Line 8 changes the cutoff for diagnosing a company that has reached the point of little likelihood of any equity payoff to the entrepreneur. The cutoff is five years in the base case; we change it to four years here. The result is to include somewhat more recent companies and to lower the certainty equivalents slightly, but not in a way that changes any of our conclusions.

Line 9 calculates the effect of the preference provisions of the standard venture contract that provide extra cash to the investors at the expense of the entrepreneurs and other common shareholders. The increase is noticeable—these provisions raise entrepreneurial risk and lower expected returns because they increase entrepreneurs’ leverage. We believe we have correctly incorporated the preferences in our calculations. This calculation illustrates the importance of the preferences, not a potential error in our work.

Lines 10 and 11 show the results of two similar alterations of our calculations. One is to boost the initial entrepreneurial ownership by one percent at the expense of the angels and other pre-venture investors. The other is to boost the ownership granted the entrepreneurs by the incentive provisions of the venture contract. Both result in a small increase in the certainty equivalent values. Line 12 shows a reverse calculation, where the share granted the nonentrepreneur employees,
such as a CEO hired to replace an entrepreneur, dilutes the entrepreneur. This alteration results in a small decline in certainty equivalent value.

Line 13 probes the sensitivity to the parameter that controls the relation between the success of the company and the entrepreneur’s incentive ownership. We chose this parameter based on our own judgment, in view of the absence of any systematic information about the incentive provisions in actual venture contracts. The calculation shows that our conclusions are not sensitive to the value of the parameter. Most entrepreneurial exit value comes from great success—IPOs or favorable acquisitions—where presumably the entrepreneurs get the maximum incentive ownership.

Line 14 drops the largest single entrepreneurial payoff, Google. It has no visible effect on the results. At a coefficient of relative risk aversion of two, the prospective entrepreneur puts very little additional weight on a billion-dollar payoff compared to a hundred-million one.

Line 15 makes the unrealistic assumption that entrepreneurs receive cash for their shares at the IPO price, at the time of the IPO, in place of our assumption that they cannot sell shares in the IPO or in the public market for six months after the IPO. The average return earned over that six-month period is over 50 percent. Nonetheless, the certainty equivalent value of the entrepreneur is higher if ownership is liquidated at the IPO and the entrepreneur forgoes the expected 50-percent return in favor of the benefit of diversification. To put it differently, the entrepreneur would be better off by selling at the low IPO price rather than taking a chance on the post-IPO appreciation (the “pop”). Typically, the IPO price drives the entrepreneur's marginal utility down to close to zero. A few companies lose all or most of their value in the six months after their IPOs. It is worthwhile for the entrepreneur to cut off this chance by selling in the IPO. Thus, the lockup period in the typical venture capital contract is another feature that imposes substantial idiosyncratic risk on the entrepreneur.

VIII. Sorting between Entrepreneurship and Employment

The coexistence of the entrepreneurial and employment contract forms for bringing new high-tech products to market presumably reflects heterogeneity on both sides of the market. Where powerful incentives are less important, large organizations will dominate because they can insure their workers. In this section, we examine sorting among individuals by deriving the crossover point for the choice of an individual between entrepreneurship and employment. We divide the three-dimensional space defined by risk aversion, alternative employment compensation, and initial assets into two subspaces, one where the individual prefers to be an entrepreneur and the complement where the individual prefers to be an employee.

Figure 10 shows the surface separating the two subspaces, as a set of lines in the risk aversion–starting assets plane. Each line shows the dividing line in the plane corresponding to a different value of the compensation available at alternative employment. The line at the lower right describes people who are indifferent between employment and entrepreneurship when employment pays a safe $300,000 per year or more, before tax. Those below and to the right of the line definitely prefer employment. The next line up and to the left shows indifference when the outside salary is $500,000. The region between the two lines describes people who are more inclined to entrepreneurship than those in the slender wedge at the bottom, because they have less risk aversion or more savings, or both. Similarly, the next line up and to the left describes indifference when outside compensation is $700,000. Those choosing entrepreneurship despite high outside earnings have quite low risk aversion or high savings.

Other characteristics may affect the sorting of engineers and scientists into entrepreneurship and employment. Those with a preference for working in an organization they help manage or a distaste for an employment hierarchy will choose entrepreneurship even if they are located
clockwise of the relevant line in Figure 10. Another possibility is that entrepreneurship attracts individuals who overestimate the likely payoffs from their ideas.

**IX. Discounting**

We stress that the parameter $r$ in our analysis is the rate of time preference and also the return earned on the entrepreneur’s savings. It is not the financial discount rate or cost of funds of the startup company. Financial discounting is implicit in the dynamic program. We can illustrate the high implicit discount by a simple thought experiment. Suppose that an entrepreneur learned in year $t$ that an exit would occur in the following year, and the entrepreneur owned a security that paid off $\epsilon X_{t+1}$ in year $t + 1$, where $\epsilon$ is a small amount. Any individual trades off small values in one period against another period at the marginal rate of substitution. Thus the value the entrepreneur would ascribe to the security would be

$$
\epsilon \frac{u'(c_{t+1})}{(1 + r)u'(c_t)} X_{t+1}.
$$

The discount factor $D_t$ is the ratio of this value to the expected value,

$$
D_t = \frac{\mathbb{E}_X \frac{u'(c_{t+1})}{(1 + r)u'(c_t)} X_{t+1}}{\mathbb{E}_X X_{t+1}}.
$$
In our base case, with coefficient of relative risk aversion of two, the discount factor $D_t$ varies from 0.01 to 0.04 over the age of the company. Thus a claim that had an expected payoff of one dollar next year, in proportion to the distribution of the exit value next year, would be worth only $0.01 to $0.04 this year. Conceptually, the discount breaks down into a pure time element and an element relating to the fact that the amount of the exit value will become known next year. The pure time discount is just the five percent in $1/(1 + r)$. All the rest of the discount comes from the uncertainty in the exit value. The reason that the entrepreneur puts such a low value on the payoff $\epsilon X_{t+1}$ is that it delivers almost all its value in circumstances where the entrepreneur is rich and has low marginal utility. Notice that $D = 1/1.05$ for a risk neutral entrepreneur with constant marginal utility.

X. Serial Entrepreneurship

Paul Gompers, Anna Kovner, Josh Lerner, and David Scharfstein (2008) report that about 12 percent of venture-backed entrepreneurs have served in that role in an earlier venture-backed startup. Our dynamic program, equation (9), does not consider that possibility. We could alter the program to include the 12 percent likelihood of future entrepreneurship, though this alteration would come at a considerable complication in calculating the value functions, because the same function would appear after the exit in the future and at time zero. The effect would be a slight increase in the value of entrepreneurship relative to employment. None of our conclusions would be significantly affected, because the probability of repeating as an entrepreneur is relatively small.

Our results have an interesting implication for serial entrepreneurship. Figure 10 shows that the choice between entrepreneurship and employment is sensitive to assets. A successful exit will give an entrepreneur a substantial level of assets, far up the vertical axis in the figure. Hence further entrepreneurship becomes far more attractive relative to employment after a success. Wealth from a successful earlier exit relieves the burden of the idiosyncratic risk of a second startup.

XI. Concluding Remarks

The contract between venture capital and entrepreneurs does essentially nothing to alleviate their financial extreme specialization in their own companies. Given the nature of the gamble revealed in Figure 2, entrepreneurs would benefit by selling some of the value that they would receive in the best outcome on the right, when they would be seriously rich, in exchange for more wealth in the most likely of zero exit value, on the left. It would be hard to find a more serious violation of the Borch-Arrow optimality condition—equality of marginal utility in all states of the world—than in the case of entrepreneurs.

A diversified investor would be happy to trade this off at a reasonable price, given that most of the risk is idiosyncratic and diversifiable. But venture capitalists will not do this—they don’t buy out startups at the early stages and they don’t let entrepreneurs pay themselves generous salaries. They use the exit value as an incentive for the entrepreneurs to perform their jobs. Moral hazard and adverse selection bar the provision of any type of insurance to entrepreneurs—they must bear the huge risk shown in Figure 2.

The venture capital institutions of the United States convert ideas into functioning businesses. We show that the process contains an important bottleneck—for good reasons based mainly on moral hazard, the venture contract cannot insure entrepreneurs against the huge idiosyncratic risk of a startup. Risk-adjusted payoffs to the entrepreneurs of startups are remarkably small. Although our results are based entirely on the venture process, we believe that no other arrangement is much better at solving the problem of getting smart people to commercialize their good ideas.
Appendix

Data

A. Investment Rounds

We use the standard and convenient vocabulary for describing the evolution of the value of a venture-backed company. When a round of funding occurs, the venture syndicate negotiates a price per share with the entrepreneurs or other management of the company. This price, multiplied by the number of shares outstanding before the new funding, is called the premoney value of the company. The sum of the premoney value and the amount newly invested is the postmoney value. The two values together fully describe the financial evolution of the company, without reference to the share prices or the number of shares. The return ratio earned by shareholders is the ratio of the new premoney value to the previous postmoney value. The premoney value is adjusted by GP fees and preferences in the case of an exit event.

Venture investors make a series of investments, \( f_1 \) through \( f_N \), in months \( t_1 \) through \( t_N \). Immediately before a round, the premoney value of the firm is \( v_i \). At time \( \tau \), the company either undergoes an initial public offering, is acquired, or ceases operations, with an exit cash payoff to the investors of \( x \).

We let \( s_{i,j} \) be the ownership share of the company attributable to the investment in round \( i \) as of round \( j \). The initial ownership share is

\[
s_{i,i} = \frac{f_i}{f_i + v_i}.
\]

Later rounds dilute the share according to the recursion,

\[
s_{i,j} = \frac{s_{i,j-1} v_j}{f_j + v_j}.
\]

The exit value of round \( i \) investors is \( x_i = s_{i,N+1} \cdot x \).

B. Adjustment of Ownership Shares in Down Rounds

A down round occurs when the share price or premoney value in one round is below the previous round. Most agreements between venture investors and entrepreneurs call for the issuance of additional shares to investors in earlier rounds when the share price in a new round falls short of the price in the previous round—what is called a “down round.” The adjustment is set forth in antidilution provisions in the agreements. Steven N. Kaplan and Per Stromberg (2003) find that about a quarter of the contracts have full-ratchet language, meaning that the entrepreneurs absorb enough of the decline in value to leave the value of venture’s ownership at the same level as in the previous round. The other three quarters of contracts have a more moderate provision called weighted average adjustment.

Both types of antidilution adjustments are modifications of equation (15) to shift ownership shares toward venture investors who paid more than the current price for their shares, where the price paid is measured on a postconversion basis. We calculate updated ownership shares for down rounds using both types of adjustment and take the weighted average, using the figures from Kaplan and Stromberg (2003).
Down-round antidilution provisions shift venture ownership upward and nonventure (entrepreneurs, angels, and employees) downward by an average of 4.8 percentage points. These provisions reduce the reward to entrepreneurs disproportionately in the less favorable outcomes.

The weighted average antidilution provision is the most common form of adjustment. To explain this provision, we let \( j \) be the number of the current round and let \( i \) range over the earlier rounds. We let \( n_{i,j} \) be the number of shares effectively held by round \( i \) investors as of round \( j \). “Effectively” means the number of common shares that would result from the investors exercising their conversion rights. Antidilution provisions take effect by lowering the conversion price, \( p_{i,j} \), and increasing the number of shares, \( n_{i,j} = f_i/p_{i,j} \), the investors receive upon conversion. We let \( N_{i,j} \) be the total number of shares outstanding at the conclusion of round \( j \).

To identify the investors eligible for the antidilution adjustment without knowing actual share counts and share prices, we proceed as follows: The conversion price per ownership share point for earlier investors as of the last round is \( f_j/s_{i,j-1} \). The price per ownership point (measured as of the previous round) paid by the new round is the premoney value \( v_j \). Thus the eligible rounds are those with \( f_j/s_{i,j-1} > v_j \). We let \( A_j \) designate the set of these rounds and \( \sim A_j \) the set of rounds not subject to adjustment, including the common shares.

The weighted average provision specifies adjustment factors for the eligible earlier investors in the case of a down round:

\[
(16) \quad a_{i,j} = \frac{N_{j-1} + n_{j,j}}{N_{j-1} + \frac{p_{j,j}}{p_{i,j-1}} n_{j,j}}.
\]

The numerator is the number of shares after round \( j \) if the existing shareholders did not receive any new shares. The denominator is the number of shares if the new round had to pay the higher price paid by an investor in round \( i \). The new conversion price is the old price divided by the adjustment factor.

The quantity \( p_{i,j} n_{j,j} \) is \( f_j \), the amount invested in the new round. The earlier conversion price \( p_{i,j-1} \) is \( f_i/n_{i,j-1} \). Thus

\[
(17) \quad a_{i,j} = \frac{N_{j-1} + n_{j,j}}{N_{j-1} + \frac{f_j}{f_i} n_{i,j-1}}.
\]

To reduce the complexity of what follows, we write

\[
(18) \quad b_{i,j} = \frac{1}{N_{j-1} + \frac{f_j}{f_i} n_{i,j-1}},
\]

so the adjustment factor is

\[
(19) \quad a_{i,j} = b_{i,j}(N_{j-1} + n_{j,j}).
\]

The total number of shares of the earlier investors, after adjustment for those who paid more than the current price, is

\[
(20) \quad \tilde{N}_j = \sum_{i \in A_j} b_{i,j}(N_{j-1} + n_{j,j})n_{i,j-1} + \sum_{i \in \sim A_j} n_{i,j-1} = B_j(N_{j-1} + n_{j,j}) + \bar{N}_j.
\]
The ownership share of the new round is

\begin{align}
    s_{j,j} &= \frac{f_j}{v_j + f_j} \\
    &= \frac{n_{j,j}}{N_j} \\
    &= \frac{n_{j,j}}{B_j(N_{j-1} + n_{j,j}) + N_j + n_{j,j}}.
\end{align}

This can be written as a linear equation in the unknown \( n_{j,j} \). We solve for \( n_{j,j} \), multiply by the expression in equation (17), and use the resulting share counts to form the new values of the ownership shares \( s_{i,j} \). By enlarging the ownership shares for the investors who paid more than the current share price, the provision reduces the shares of the entrepreneurs and other earlier investors even more than the normal dilution from a new round.

The calculations described above are homogeneous in the numbers of shares, so we can normalize the total number of shares from the previous round at one. After this normalization, \( n_{i,j-1} = s_{i,j-1} \). The effect of the calculations, including forming the new shares \( s_{i,j} \), is to modify equation (15) to include the rearrangement of equity interests among the existing shareholders that occurs in a down round.

The description of the weighted average updating in a down round given in this Appendix is rather more complicated than in standard references on venture contract terms, such as Bagley and Dauchy (2003). Those descriptions assume the availability of data on share holdings and conversion prices. Our approach is tailored to our data, which require us to infer these numbers from data on pre- and postmoney value.

In the case of the full-ratchet antidilution adjustment, the rearrangement of ownership shares can be expressed in the same framework. Those investors who paid more than the current price for their shares in an earlier round receive a proportional increase in ownership (decrease in conversion price) equal to the ratio of the earlier price to the current price. If an earlier round, \( i \), had a higher price, its number of shares becomes \( f_i/p_{j,j} \). As before, the current price is \( p_{j,j} = f_j/n_{j,j} \). Thus the number of adjusted shares brought into the current round is

\begin{equation}
    \sum_{i \in A_j} \frac{f_i}{f_j} n_{j,j} + \sum_{i \notin A_j} n_{i,j-1}.
\end{equation}

Again, we can solve the equation for the ownership share of the new round,

\begin{equation}
    s_{j,j} = \frac{n_{j,j}}{n_{j,j} \sum_{i \in A_j} \frac{f_i}{f_j} + \sum_{i \notin A_j} n_{i,j-1} + n_{j,j}},
\end{equation}

for the new ownership \( n_{j,j} \), and then calculate the ownership shares of the earlier investors and entrepreneurs. As before, the total number of shares owned as of the previous round can be normalized at one, so the procedure developed here is a recursion that describes the rearrangement in ownership shares that occurs in a down round because of the antidilution provision benefiting earlier venture investors.
Sources

We use a database compiled by Sand Hill Econometrics on venture investments in startups and on the fates of venture-backed companies. The data are drawn from a variety of sources, including several commercial data vendors. The vendors concentrate on reporting funding and valuations for venture investments and are less likely to report exit events, especially shutdowns and acquisitions at low values. Sand Hill Econometrics has used a wide range of sources to augment coverage of these adverse termination events.

One important source of valuation data is S-1 statements filed by venture-backed companies when they go public. These statements often give a funding history for the company. Because an IPO is a favorable event, the back-filling of round values from S-1s is a source of return based selection in the data. Table A1 describes the data. Our general database reports 62,609 funding rounds for 22,004 companies. Among the exit values used in the analysis, 2,015 are IPOs, 5,625 are acquisitions, and 3,352 are confirmed zero value exits. For an additional 4,220 companies, we infer zero value exits from the observation that the company neither exited nor raised funds in the last five years of our sample. We assign an exit date to these companies by drawing from the distribution of time from last funding to exit for the companies with known failure dates.

Of the 62,609 funding rounds included in the analysis, we can infer the venture share directly from the reported value in 16,637 of the rounds. In the remaining 45,972 rounds, we impute the venture share of ownership as described below. For this purpose, we use the second look database of 1,292 funding rounds where the values (and thus venture shares) are reported for companies with missing valuations in the general database. That subsample contains 762 rounds.

Preferences

The standard financial contract gives venture convertible preferred equity in a company—see Kaplan and Stromberg (2003), Table 1. In about half of these cases, venture receives both its original investment and its common stock value, sometimes with an upper limit on the common stock value for payment of the original investment. The second form of claim is called participating preferred stock.

We do not observe the preference terms for each of the rounds of investment in our database. Further, we observe unfavorable outcomes—those in the range where the preferences would
matter—after negotiations that altered the preferences may have occurred among the disappointed claimants on a company. We understand that venture investors sometimes bargain away their preferences and become common shareholders in order to induce the entrepreneurs to agree to a disappointing exit plan. Wilmerding (2003) writes, in connection with a low potential exit value, “...at those levels, where management will not receive much from a sale, the preferred shareholders will likely be forced to give up some of their return to make the deal work” (p. 52). Our data do not reveal if the cash from a low value exit is distributed according to the original contracts or whether the parties have bargained to a jointly superior outcome once the bad news arrived.

The reason that a jointly superior bargain is available is that adverse events leave the entrepreneurs, holding only common stock, with option positions that are far out of the money because of the preferences. In this situation, the entrepreneurs have little incentive to perform the functions needed to recover limited value from a disappointing outcome. They will prefer to continue rolling the dice unless a new deal can be struck that better aligns incentives.

In some cases, one round of preferred stock has priority over another, but we lack information on priorities, so we assume that all rounds of preferred shares have equal priority and divide the available cash in proportion to the amount invested. Kaplan and Stromberg (2003), Table 2, report that 71 percent of venture contracts grant preferences in excess of the amount invested, often in the form of a cumulative dividend, which averages 8 percent per year. We take the preference to be 125 percent of the amount invested, corresponding to about 3 years of the dividend.

For nonparticipating preferred shares, the cash payout is

\[
\max(\min(\tilde{s}X, \tilde{s}P), sX),
\]

where \(\tilde{s}\) is the share of a given venture round among all venture (preferred) shareholders and \(s\) is the share of the round among all shareholders, \(X\) is the exit value of the company, and \(P\) is the preference amount for the round. For participating shares, the payout is

\[
\min(\tilde{s}X, \tilde{s}P + s(X - P)).
\]

E. Additional Ownership for Entrepreneurs and Employees

The contract between venture investors and entrepreneurs often includes provisions for additional ownership based on company performance. The entrepreneurs vest in the shares upon reaching milestones in the contract. In addition, nonfounder employees vest in stock options based on longevity and other factors. Kaplan and Stromberg (2003), Table 2, report on both elements of vesting in terms of ownership shares if no vesting occurs and if all vesting occurs. The average founders’ share rises from 24.3 percent to 31.1 percent upon full vesting and the average share of nonfounder and employee ownership rises from 20.2 percent to 22.2 percent.

We interpret the initial 20.2 percent as the share of nonentrepreneur, nonventure investors—angels and friends of the family. We interpret the initial nonventure share observed in our data as the sum of the entrepreneurs’ share and the 20.2 percent. We interpret the incremental 2.0 percent as the ownership of nonfounder employees upon a successful exit such as an IPO. We interpret the incremental 6.8 percent of entrepreneur ownership in the same way.

We construct an index of success, with respect to vesting, as

\[
z = \min\left(1, \frac{X}{\psi F}\right),
\]
where \( X \) is the gross value to all existing shareholders at exit and \( F \) is the sum of all rounds of venture investing. The parameter \( \psi \) is interpreted as the ratio where all possible vesting occurs. We take \( \psi = 8 \). We boost the entrepreneur and nonventure investor share by 0.068 \( z \) and impute a share to nonfounder employees of 0.02 \( z \).

### F. Imputing Missing Venture Ownership Shares

In our data, the amounts invested by venture are reported quite fully. As Table A1 shows, however, the valuation of the company at the time of the investment is often not reported. We impute the missing data on the form of the ownership share acquired through the new investment. The share implies pre- and postmoney values of the company at the time of the investment.

We impute missing data for the venture shares by combining a standard missing data approach with a unique body of data that provides a full solution to the problem of selection bias that plagues the imputation of missing data in most applications.

Our second-look database gives full information about valuations obtained from another source for more than a thousand of the financing rounds in the general data. We make the imputations of venture share from the second look database. Here we know about missing data, from the perspective of the large body of data, and about true shares, from the fully reported data. Thus we can make a direct attack on the selection problem described above. We fit an equation to the actual shares in the second look data for the companies with missing valuations in the general database.

The second-look database contains data on 2.0 percent of the funding rounds in our main database. It is reasonably representative of the main database. We measure representativeness as the ratio of funding rounds in a given category to the expected number if the second-look data were perfectly representative and thus contained 2.0 percent of the number of rounds in that category in the main database. The second-look data were collected in 2006 and do not contain any rounds since then; they also lack rounds from before 1993. The representativeness ratio is 96 percent for 1993 to 1995, 117 percent for 1996 to 1998, 155 percent for 1999 to 2001, and 149 percent for 2002 to 2004. Among major sectors, the ratio is 59 percent for information technology, 104 percent for telecommunications, 162 percent for biotechnology, and 146 percent for retail. The ratio is 150 percent for companies that eventually went public. For rounds reported in the second look data, the fraction reporting value in the main data was 31 percent, compared to 24 percent for all rounds in the main data.

The company values reported in the second-look data (which we regard as virtually certain to be correct in the sense of being the values actually calculated by the venture funds) are generally identical to or close to the values reported in the main database. About 60 percent of the 402 rounds present in the second-look data for which values are also reported in the main data have exactly the same value in both sources. About 90 percent of the overlap values in the main database are within 15 percent of the second look values. The maximum discrepancy is 41 percent.

The venture share, \( s \), needs to be tracked through the various rounds of financing as later rounds dilute the ownership of earlier rounds. The calculations require the most recent share,

\[
(27) \quad s_{i,i} = \frac{f_i}{f_i + v_i}
\]

because the recursion in equation (15) can be written:

\[
(28) \quad s_{i,j} = s_{i,j-1}(1-s_{i,i})
\]
We use the nonlinear logit regression specification

\[ s_{i,t} = \frac{1}{1 - \exp(-X_i \delta)}, \]

where \( X_i \) is a vector of variables observed even when \( v_i \) is missing and \( \delta \) is the corresponding vector of parameters.

As predictive variables, we use:

- number of this round
- amount raised in this round
- cumulative increase in the Wilshire index over the two years preceding this round.

Our specification has a complete set of interactions by round number, except that we fit the same coefficients for rounds 5 and higher.

Table A2 shows the results of this regression.

The selection bias for the main data appears to be downward: The average value for rounds reported in both the main data and the second look data is $22 million less than the average value reported in the second look data for rounds in the main data where value was not reported. The difference is statistically unambiguous, with a standard error of $7 million.

### Joint Distribution with Positive Dependence of the Entrepreneur’s Exit Value and Share of That Value

We define the latent random variable \( z \) related to total entrepreneurial exit value \( V \) by

\[ z = \rho V + u. \]
Here $u$ is a random variable distributed uniformly on the unit interval. The parameter $\rho$ controls the strength of the dependence. The support of $z$ is $[\rho V, \rho V + 1]$, where $[V, \bar{V}]$ is the range of values of $V$ in the database. The marginal cdf of $z$ is

$$\text{Prob}[z \leq Z] = \sum_{V} \text{Prob}[z \leq Z | V] \text{Prob}[V],$$

where the summation is over all the values of $V$ in the database. $\text{Prob}[V]$ is the reciprocal of the number of exit values $V$ in the database. The conditional probability is

$$\text{Prob}[z \leq Z | V] = \text{Prob}[u \leq Z - \rho V] = \begin{cases} Z - \rho V & \text{if } Z - \rho V \leq 0 \\ 1 & \text{if } Z - \rho V \geq 1 \end{cases} = \begin{cases} Z - \rho V & \text{if } 0 \leq Z - \rho V \leq 1 \\ 0 & \text{otherwise} \end{cases}.$$
We evaluate the conditional probability at 1,000 values of \( Z \), using all of the values of \( V \) in the database. We then find the cutoff values of \( Z_i \), corresponding to the distribution of the entrepreneur’s share among the six categories. Thus, 

\[
\text{Prob} \left[ s \in S_i \right] = \text{Prob} \left[ z \in [Z_{i-1}, Z_i] \right].
\]

Here \( S_i \) is the range of values of the entrepreneur’s share shown in the first column of text Table 1.

When solving the consumer dynamic program, (9), we form the expectation over the entrepreneur’s exit value, \( X = sV \), by summing over the \( V \)s in a given age group and the conditional distribution of the share, \( s \), in the six categories, each with probability 

\[
\text{Prob} s = \bar{s}_i = \text{Prob} \left[ z \leq Z_i \mid V \right] - \text{Prob} \left[ z \leq Z_{i-1} \mid V \right],
\]

where equation (32) gives the conditional probability. The values \( \bar{s}_i \) are the means of the share within each category, shown in the right-most column of text Table 1.

We calculated the expected company exit values given the entrepreneur’s share category as follows. The probability that a given company exit value \( V \) corresponds to entrepreneur share category \( i \) as defined in text Table 1 is 

\[
\min \left( 1, Z_i - \rho V \right) - \max \left( 0, Z_{i-1} - \rho V \right).
\]

We calculated \( \mathbb{E} (V \mid i) \) as the sum of the \( V \)s in the sample weighted by these probabilities. We found that \( \rho = 0.011 \) equated \( \mathbb{E} (V \mid i = 6) / \mathbb{E} (V \mid i = 1) \) to its value in the IPO sample of 1.9.

Table A3 shows the joint distribution of entrepreneurial exit value and venture lifetime corresponding to text Table 2, but with the alternative assumption of correlation between total entrepreneurial exit value and an entrepreneur’s share. We calculate the distribution by using the conditional probabilities in (35) in place of \( h_i \) in text (3). Note that the marginal distributions of the entrepreneur’s exit value are slightly different in Table A3 from text Table 2, because we do the calculation prior to renormalizing the joint distribution.

### Bootstrap Standard Errors

The estimates in Table 3 have sampling variation from two sources: (i) estimation of the coefficients of the equation for imputing shares when company value is not reported, and (ii) the
use of a finite actual sample in the dynamic program in place of the distribution from which the sample is drawn. To estimate the resulting sampling variation, we use a Monte Carlo approach for the first source and bootstrap for the second. Our procedure makes the reasonable assumption that the sampling distribution of the coefficients is independent of the draws for the dynamic program, because the second-look data constitute a tiny fraction (two percent) of all the companies included in the dynamic program.

For the sampling distribution of the coefficients, we drew from a multivariate normal distribution with the reported covariance matrix from the estimation of the coefficients. For the bootstrap, we drew samples with replacement from the tabulated exits. We drew 25 samples, each combining a Monte Carlo version of the coefficients and a bootstrap set of exit values. We then performed the dynamic program calculations underlying Table 3. We report the standard deviation of the samples in Table A4. These calculations take about two days using a vintage-2008 personal computer and Matlab.

REFERENCES


